

Calibration Concordance

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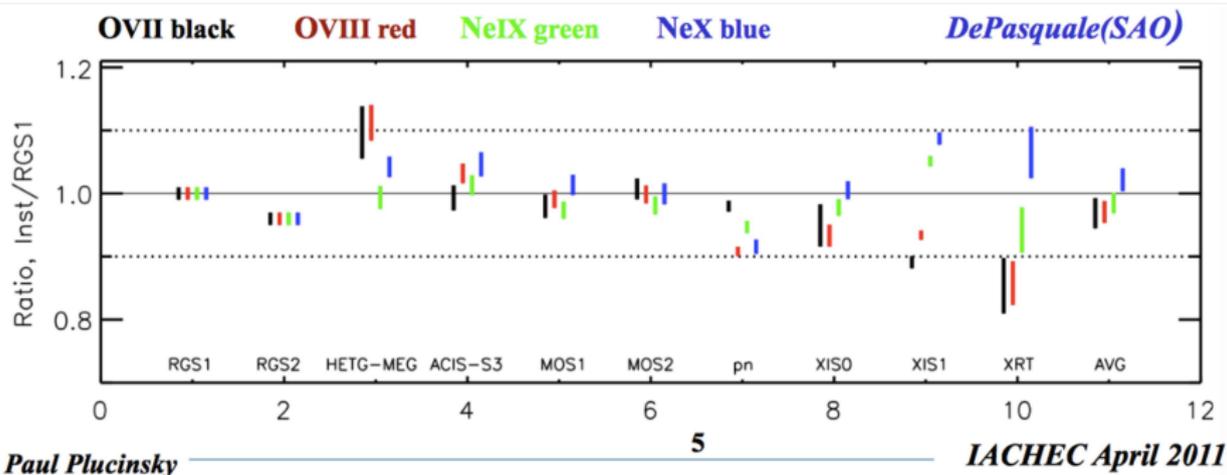
August 21, 2017

Calibration Concordance Problem (Example: E0102)



- Supernova remnant E0102
- Four sources correspond to four spectral lines in E0102

Calibration Concordance Problem (Example: E0102)



- Four spectral lines observed with 11 X-ray detectors
- Main challenge – the data/instruments do not agree

Outline

- 1 Introduction
- 2 Scientific and Statistical Models
- 3 Bayesian Hierarchical Model
- 4 Advantages of Our Approach
- 5 Summary

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 - For each source j , F_j is unknown.
- Photon counts c_{ij} : from measuring flux F_j with instrument i .
- Lower cases: data / estimators.
- Upper cases: parameter / estimand.

Calibration Concordance Problem

- 1 Astronomers' Dilemma:

$$\frac{c_{ij}}{a_i} \neq \frac{c_{i'j}}{a_{i'}} \text{ for } i \neq i'.$$

Different instruments give different estimated flux of the same object!

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2 Scientific Question:

- Are there systematic errors in 'known' effective areas?
- Can we derive properly adjusted effective areas?
- Can we unify estimates of the same flux with different instruments?

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Scientific and Statistical Models

Scientific Model

Multiplicative in original scale and additive on the log scale.

$$\text{Counts} = \text{Exposure} \times \text{Effective Area} \times \text{Flux},$$

$$C_{ij} = T_{ij}A_iF_j, \quad \Leftrightarrow \quad \log C_{ij} = B_i + G_j,$$

where $\log \text{area} = B_i = \log A_i$, $\log \text{flux} = G_j = \log F_j$; let $T_{ij} = 1$.

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Statistical Model

$$\log \text{ counts } y_{ij} = \log c_{ij} - \alpha_{ij} = B_i + G_j + e_{ij}, \quad e_{ij} \stackrel{\text{indep}}{\sim} \mathcal{N}(0, \sigma_{ij}^2);$$

where $\alpha_{ij} = -0.5\sigma_{ij}^2$ to ensure $E(c_{ij}) = C_{ij} = A_iF_j$.

- **Known Variances:** σ_{ij} known.
- **Unknown Variances:** $\sigma_{ij} = \sigma_i$ unknown.

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Bayesian Hierarchical Model

Log-Normal Hierarchical Model.

$$\text{log counts} \mid \text{area \& flux \& variance} \stackrel{\text{indep}}{\sim} \text{Gaussian distribution,}$$

$$y_{ij} \mid B_i, G_j, \sigma_i^2 \stackrel{\text{indep}}{\sim} \mathcal{N}(B_i + G_j, \sigma_i^2),$$

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 B_i & \stackrel{\text{indep}}{\sim} \mathcal{N}(b_i, \tau_i^2), \\
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 \text{If variance unknown: } \sigma_i^2 & \stackrel{\text{indep}}{\sim} \text{Inv-Gamma}(df_g, \beta_g).
 \end{aligned}$$

Setting the prior parameters.

- 1 $b_i = \log a_i$, τ_i are given by astronomers.

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log counts | *area & flux & variance* $\overset{\text{indep}}{\sim}$ Gaussian distribution,

$$y_{ij} | B_i, G_j, \sigma_i^2 \overset{\text{indep}}{\sim} \mathcal{N}(B_i + G_j, \sigma_i^2),$$

$$B_i \overset{\text{indep}}{\sim} N(b_i, \tau_i^2),$$

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Setting the prior parameters.

- ① $b_i = \log a_i, \tau_i$ are given by astronomers.
- ② df_g, β_g are given based on the variability in data.

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Advantages of Our Approach

- ① Intuitive Interpretation: Shrinkage Estimators
- ② Adjusted Estimates of Effective Area
- ③ Calibration Concordance
- ④ Avoiding Pitfalls of Wrong 'Known Variances'

Shrinkage Estimators: Known Fluxes and Errors

Hierarchical model \Rightarrow Shrinkage estimators

- weighted averages of evidence from 'Prior' and evidence from 'Data').

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When fluxes and variances are known,

Original Scale

$$\hat{A}_i = a_i^{W_i} \left[(\tilde{c}_{i\cdot} \tilde{f}^{-1}) e^{\sigma_i^2/2} \right]^{1-W_i},$$

where

$$\tilde{c}_{i\cdot} = \prod_j c_{ij}^{1/M}, \quad \tilde{f} = \prod_j f_j^{1/M}$$

are geometric means.

Log-Scale

$$\hat{B}_i = W_i b_i + (1 - W_i)(\bar{y}_{i\cdot} - \bar{G}),$$

where

$$\bar{G} = \frac{\sum_j g_j}{M}, \quad \bar{y}_{i\cdot} = \frac{\sum_j y_{ij}}{M}$$

are arithmetic means.

The 'weights', $W_i = \frac{\tau_i^{-2}}{\tau_i^{-2} + M\sigma_i^{-2}}$, represents the direct information in b_i relative to indirect information in fluxes.

Shrinkage Estimators: Known Errors

When fluxes are unknown and variances are known,

$$\hat{B}_i = W_i b_i + (1 - W_i)(\bar{y}_{i\cdot} - \bar{G}_i), \quad \hat{G}_j = \bar{y}_{\cdot j} - \bar{B},$$

$$\text{where } \bar{G}_i = \frac{\sum_j \hat{G}_j}{M}, \quad \bar{B} = \frac{\sum_i \hat{B}_i \sigma_i^{-2}}{\sum_i \sigma_i^{-2}}, \quad \bar{y}_{i\cdot} = \frac{\sum_j y_{ij}}{M}, \quad \bar{y}_{\cdot j} = \frac{\sum_i y_{ij} \sigma_i^{-2}}{\sum_i \sigma_i^{-2}}.$$

In practice, we use MCMC to fit the full model.

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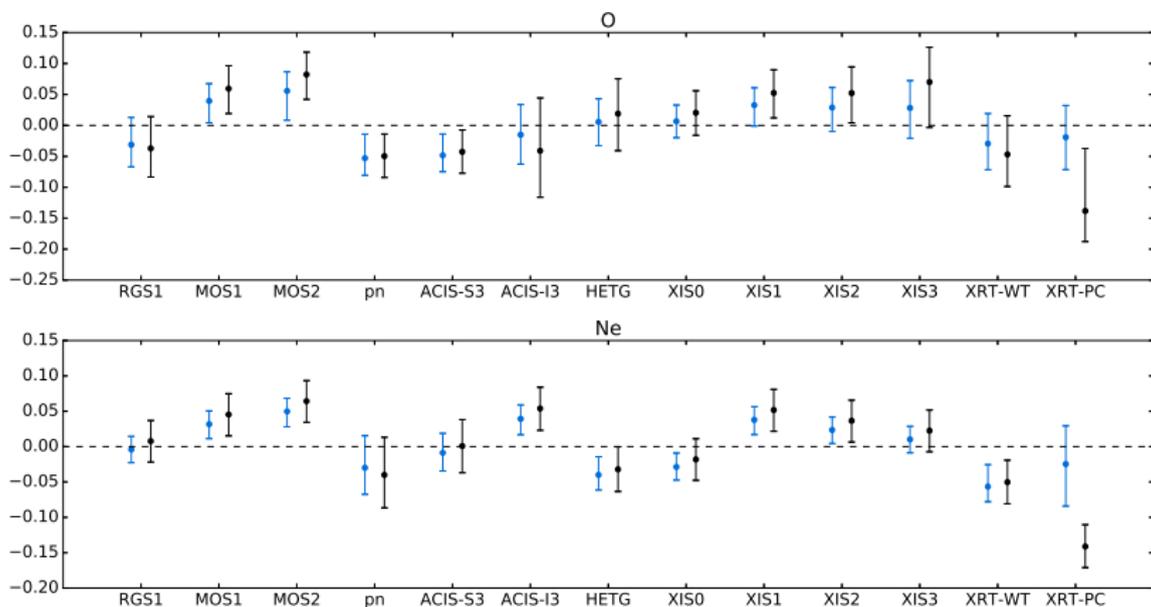
Numerical Results (E0102)

Recap: Supernova remnant E0102.

Four sources are four spectral lines in E0102.



Estimates of $B_i = \log A_i$ ($M = 2$ each panel)

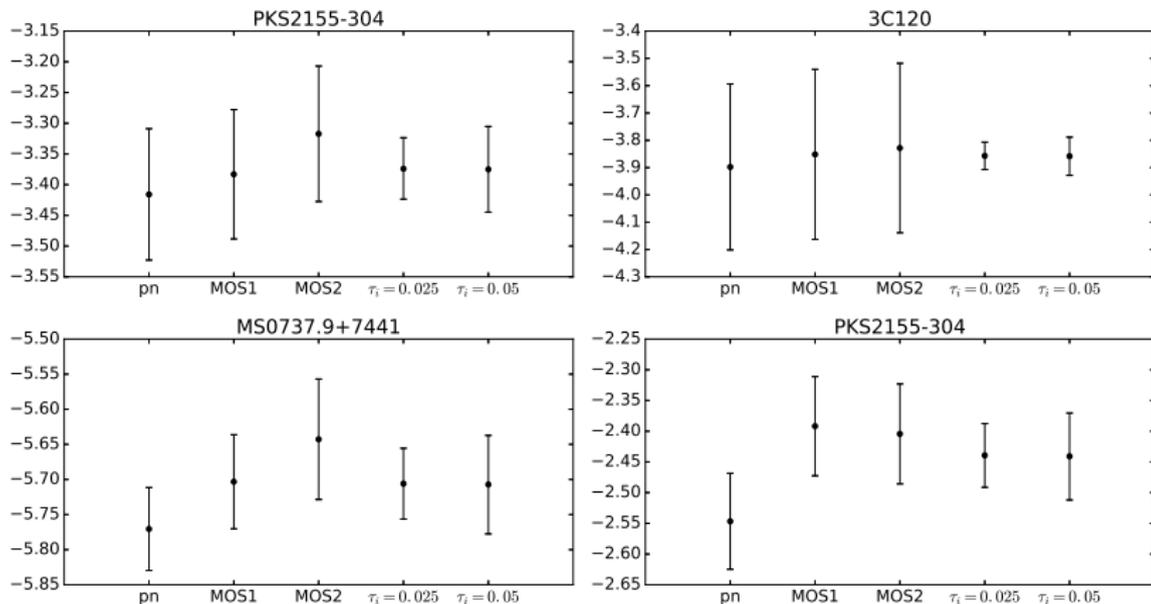


- Adjusted so that default effective area, $b_i = \log a_i = 0$.
- 95% posterior intervals (black: $\tau = 0.05$; blue: $\tau = 0.025$).
- Some instruments systematically high, others low.

Numerical Results (XCAL)

- **XCAL data:** Bright active galactic nuclei from the XMM-Newton cross-calibration sample.
- **Pileup:** Image data are clipped to eliminate the regions affected by pileup, determined using `epatplot`.
- **Three detectors:** MOS1, MOS2 and pn.
- **Sources:** $M=103$ (in medium band).

Numerical Results (XCAL): Calibration Concordance



- y-axis: G (log flux)
- vertical bars (left 3): mean \pm 2 s.d. based on observed fluxes
(right 2): 95% our posterior intervals.
- Calibration Concordance: *A single estimate of each flux!*

Benefits of Fitting σ_i^2

- Tolerance to model/error model misspecification.

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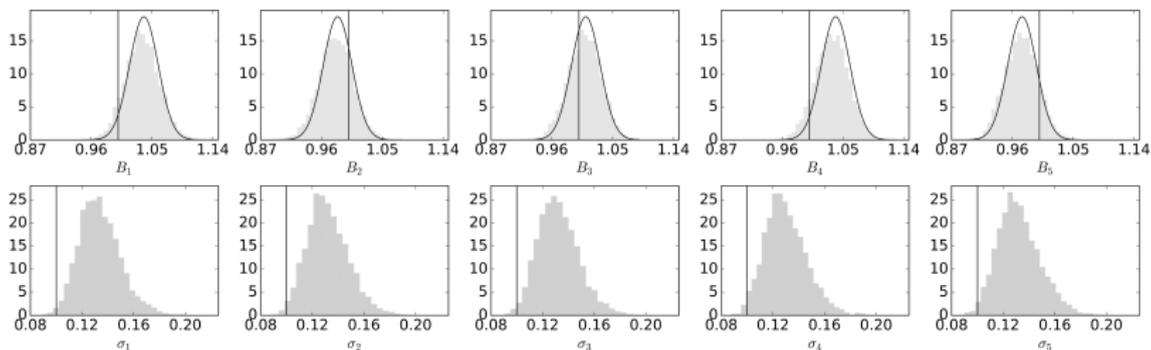
- Tolerance to model/error model misspecification.
- Pitfalls of assuming 'known' variances:
 - Overly optimistic 'known variances'
 - ⇒ overly narrow confidence intervals
 - ⇒ possible false discoveries

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- Tolerance to model/error model misspecification.
- Pitfalls of assuming 'known' variances:
 - Overly optimistic 'known variances'
 - ⇒ overly narrow confidence intervals
 - ⇒ possible false discoveries
 - 'known variances' \geq true variability
 - ⇒ noninformative results

Benefits of Fitting σ_i^2 : Example

Simulated Data: Poisson data with $N = 10$, $M = 40$, $B_i = 1$, $G_j = 3$.



Histograms: posterior distributions.

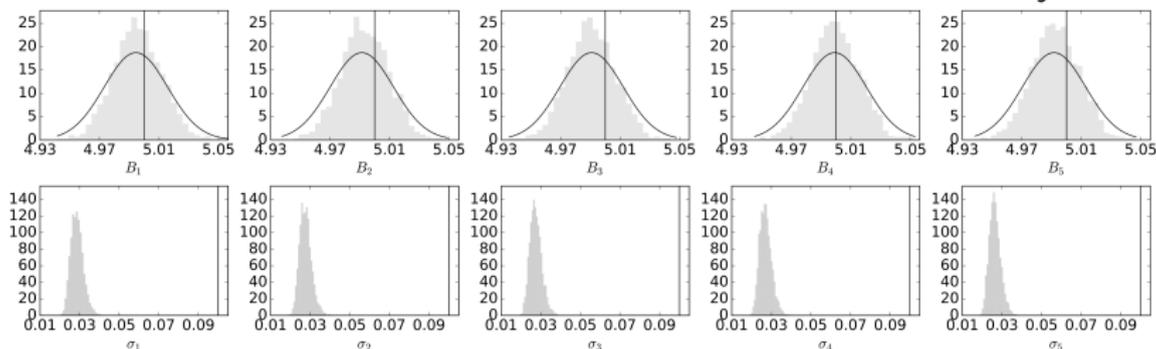
Vertical line: true values

Black Curve: Results with 'known variances' $\sigma_i^2 = 0.1^2$, (\approx fit)

No cost to fitting σ , even when values are known correctly.

Benefits of Fitting σ_i^2 : Example

Simulated Data: Poisson data with $N = 10$, $M = 40$, $B_i = 5$, $G_j = 3$.



Histograms: posterior distributions.

Vertical line: true values

Black Curve: Results with 'known variances' $\sigma_i^2 = 0.1^2$, ($>$ fit)

When 'known' σ is off, under/over estimate errors in fit.

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- 1 Adjustments of effective areas of each instrument.

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Astronomy

- 1 Adjustments of effective areas of each instrument.
- 2 Calibration concordance.

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