# A Statistical Approach to Stellar Archaeology 

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Aug. 1, 2007


## Color-Magnitude Diagram (CMD)

The CMD of our target data set.
Color Magnitude Diagram of NGC 346


## Isochrones - Reference Tables

Isochrones reveal physical informations from CMD (Temp, mass, [ $\mathrm{Fe} / \mathrm{H}]$, age, class, etc.)

- Realization of theoretical models (quite complicated).
- Depends on so many input parameters.
- Tables tell the location of a star in the diagram (color vs magnitude) at the given age, metalicity, mass, etc.
- A typical inverse problem in astronomy

How these isochrones look like on CMD?

## Isochrones continued- How they look?

Age 3000 yr to 15.5 Gyr isochrones, $[\mathrm{Fe} / \mathrm{H}]=0.4$


100Myr isochorones with different [Fe/H]


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## 47 Tuc: 11.2 Gyr Old Globular Cluster

NGC 104 (47 Tuc)


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## Motivation

CMD (data) and Isochrones (model) are available and we like to know the age distribution of stellar clusters in a statistical fashion.

## Statistical Modeling- Simplest

Bayes Rule: the age $(\tau)$ posterior distribution is proportional to the likelihood times a prior. For a single star,

$$
p\left(\tau \mid M_{i}, C_{i}\right) \propto l_{i}\left(M_{i}, C_{i} \mid \tau\right) p(\tau)
$$

where $M_{i}$ is magnitude $\left(M_{v}\right)$ and $C_{i}$ color $(B-V)$. Therefore,

$$
p\left(\tau \mid\left\{M_{i}\right\},\left\{C_{i}\right\}\right) \propto p^{n}(\tau) \prod_{i=1}^{n} l_{i}\left(M_{i}, C_{i} \mid \tau\right)
$$

We like to focus on the likelihood,

$$
\prod_{i=1}^{n} l_{i}\left(M_{i}, C_{i} \mid \tau\right)
$$

How can we estimate this likelihood?

## Information Theory: K-L distance

The Kullback-Leibler distance is defined to be

$$
D\left(f_{\tau}(x) ; g(x \mid \tau)\right)=\int \log f_{\tau}(x) f_{\tau}(x) d x-\int \log g(x \mid \tau) f_{\tau}(x) d x \geq 0
$$

We do not know the true age density $f_{\tau}(x)$ but introducing $g(x \mid \tau)$ and maximizing $E_{\tau}[(\log g(X \mid \tau)]$ provides the best $\tau$ (age) for a given $g(x \mid \tau)$, where $x$ is observed and its random variable is denoted by $X$.

$$
\hat{\tau}=\arg \max _{\tau \in\left\{\tau_{j}\right\}} E_{\tau}[\log g(X \mid \tau)]
$$

To estimate $E_{\tau}[\log g(X \mid \tau)]$, we used the empirical mean of the $\log$ likelihood.

$$
E_{\tau}[\log g(X \mid \tau)]=\frac{1}{n} \sum_{i=1}^{n} \log g\left(x_{i} \mid \tau\right)+\frac{b}{n}
$$

where $b$ is a bias term such as the penalty terms in AIC and BIC. We assume $\frac{b}{n} \rightarrow 0$. Finding $\tau$ that maximizes $\frac{1}{n} \sum_{i=1}^{n} \log g\left(x_{i} \mid \tau\right)$ leads to the best guess for the age of the stellar cluster.

Then, what would be the likelihood?

## Likelihood

We took Multivariate Normal to establish the likelihood based on additional informations from data:

- Errors on each observation ( $\sigma_{i}$ are known)
- Independence among color bands (U,B,V,I,R,etc)
- Multivariate Normal assumption is quite reasonable For a star $i$, the recorded value is apparent magnitudes $\left(v_{i}, b_{i}\right)$, corrected by the appropriate distance modulus and the extinction laws (these corrections have their own uncertainties but we ignore at the moment).


## Likelihood - continued

By denoting this corrected value as $\left(M_{i}, C_{i}\right)=x_{i}$, the likelihood of a star is

$$
l_{i}\left(M_{i}, C_{i} \mid \tau\right)=\frac{1}{2 \pi\left|\Sigma_{i}\right|^{1 / 2}} \exp \left(-\frac{1}{2}\left(x_{i}-\mu_{i}\right)^{T} \Sigma_{i}^{-1}\left(x_{i}-\mu_{i}\right)\right)
$$

where $\Sigma_{i}$ is a covariance matrix for star $i$ and $\mu_{i}=\mu_{i}(\tau,[\mathrm{Fe} / \mathrm{H}]$, mass, class $)$ although the function $\mu_{i}$ is unknown. As indicated, the best tactic is maximizing the likelihood and this is achieved by finding $\mu_{i}$ that minimize the distance to $x_{i}$. Given $\tau_{j}\left(j=1, \ldots, 91\right.$, Geneva models provide 91 age grids), $\mu_{i}\left(h_{j}\right)$ minimizes the distance to $x_{i}$.

But how to minimize the distance?

## Defining a point of min. distance

Finding a point of minimum distance, associated with a complicated curve, only represented by a set of points. $\rightarrow$ piece-wise Euclidean distance


## Applying the model selection method

Real Data Application:

| NGC | $[\mathrm{Fe} / \mathrm{H}]$ | $(m-M)_{o}$ | $E(B-V)$ | Age $^{1}(\mathrm{Gyr})$ | Est. Age |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 104 (Tuc47) | 0.004 | 13.33 | 0.04 | $10.9 \pm 1.4$ | 11.22 |
| 188 | 0.02 | 11.17 | 0.09 | $6.3 \pm 0.8$ | 5.0 |
| 2420 | 0.007 | 11.94 | 0.05 | $2.2 \pm 0.3$ | 2.24 |
| $2682(\mathrm{M} 67)$ | 0.02 | 9.59 | 0.04 | $4.3 \pm 0.5$ | 3.55 |
| 6791 | 0.050 | 12.96 | 0.15 | $10.2 \pm 1.2$ | 8.71 |
| 7789 | 0.014 | 11.22 | 0.29 | $1.80 \pm 0.3$ | 2.24 |

## Log Likelihood Profiles of Some Stellar Clusters

likelihood profiles


## Age of NGC 346 - part of SMC

Isochrones (3000-15.5G yrs), [Fe/H]=0.01


## Log Likelihood Profiles of NGC 346 and SMC

## likelihood profiles

- NGC346 (17.8Myr)
- SMC (49.0Myr)



## Discussion

The model selection by maximizing likelihoods is

- empirical method: the correctness of ages highly depends on data. (data processing affects results: e.g. foreground stars and covariance matrix).
- easy, quick, simple, heuristic, and diagnostic.

Requires a fine tuning: Extention to Bayesian approaches

- posterior distribution from marginalizing the initial mass function (IMF)

$$
\begin{aligned}
p\left(\tau \mid\left\{M_{i}\right\},\left\{C_{i}\right\}\right) & =\int p\left(\tau, m \mid\left\{M_{i}\right\},\left\{C_{i}\right\}\right) d m \\
& \propto \int \prod_{i=1}^{n} l_{i}\left(M_{i}, C_{i} \mid \tau, m\right) p(\tau, m) d m
\end{aligned}
$$

- developing hierarchical models to incorporate not only IMF but metalicity, classes, completeness, and uncertainties, etc...

Thank you!

Please visit AstroStatistics blog at http://groundtruth.info/AstroStat/slog/

